Statistics Preliminaries

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The Normal Distribution

 \bullet We say that a random variable (RV) X is normally distributed with mean μ and standard deviation σ if its probability density function is:

$$
f(x) = \frac{1}{\sqrt{2\pi\sigma^2}}e^{-\frac{(x-\mu)^2}{2\sigma^2}}
$$

and we usually write $X \sim \mathsf{N}(\mu, \sigma^2).$

Standard Normal Cumulative Probability

• If $Z \sim N(0, 1)$, we have that:

Since the integral cannot be solved in closed-form, the probability must then be obtained from a table or using a computer.

Computing Normal Probabilities

In order to compute cumulative probabilities for X ∼ N(*µ, σ*²), we usually normalize the random variable X :

$$
\Pr(X \le x) = \Pr\left(\frac{X-\mu}{\sigma} \le \frac{x-\mu}{\sigma}\right) = \Pr\left(Z \le \frac{x-\mu}{\sigma}\right) = \Phi\left(\frac{x-\mu}{\sigma}\right)
$$

where $Z = \frac{X-\mu}{\sigma} \sim \mathsf{N}(0,1)$ is called a Z-score.

Example 1

Suppose that $X \, \sim \, {\sf N}(\mu,\sigma^2)$ with $\mu \, = \, 10$ and $\sigma \, = \, 25.$ What is the probability that $X \leq 0$?

$$
Pr(X \le 0) = Pr\left(Z \le \frac{0-10}{25}\right) = \Phi(-0.40) = 0.3446.
$$

Computing a Right-Tail Probability

• Right-Tail Probability: $Pr(X > x) = 1 - Pr(X \leq x)$

Example 2

Suppose that $X \, \sim \, {\sf N}(\mu,\sigma^2)$ with $\mu \, = \, 10$ and $\sigma \, = \, 25.$ What is the probability that $X > 12$?

$$
Pr(X \le 12) = Pr\left(Z \le \frac{12-10}{25}\right) = \Phi(0.08) = 0.5319.
$$

Therefore, $Pr(X > 12) = 1 - 0.5319 = 0.4681$.

Computing an Interval Probability

 \bullet Interval Probability: $Pr(x_1 < X \leq x_2) = Pr(X \leq x_2) - Pr(X \leq x_1)$

Example 3

Suppose that $X \, \sim \, {\sf N}(\mu,\sigma^2)$ with $\mu \, = \, 10$ and $\sigma \, = \, 25.$ What is the probability that $2 < X < 14$?

$$
Pr(X \le 14) = Pr\left(Z \le \frac{14-10}{25}\right) = \Phi(0.16) = 0.5636,
$$

Pr(X \le 2) = Pr\left(Z \le \frac{2-10}{25}\right) = \Phi(-0.32) = 0.3745.

Therefore, $Pr(2 < X < 14) = 0.5636 - 0.3745 = 0.1891$.

Computing a Standard Normal Right-Tail Percentile

 \bullet For a standard normal variable Z, a right-tail percentile is the value z*^α* above which we obtain a certain probability *α*.

• Mathematically, this means finding z_α such that:

$$
\Pr(Z > z_\alpha) = \alpha \Leftrightarrow \Pr(Z \leq z_\alpha) = 1 - \alpha
$$

The following table shows common values for z*α*:

Confidence Interval for a Standard Normal RV

 \bullet A 1 – α confidence interval (CI) defines left and right percentiles such that the probability on each side is *α/*2.

 \bullet For a standard normal variable Z, the symmetry of its PDF implies: $\mathsf{Pr}(Z \leq -z_{\alpha/2}) = \mathsf{Pr}(Z > z_{\alpha/2}) = \alpha/2$

Example 4

Since $z_{2.5\%} = 1.96$, the 95% confidence interval of Z is $[-1.96, 1.96]$. This means that if we randomly sample this variable 100,000 times, approximately 95,000 observations will fall inside this interval.

Confidence Interval for a Normal RV

If X ∼ N(*µ, σ*²), its confidence interval is determined by *ξ* and *ζ* such that:

$$
\Pr(X \le \xi) = \alpha/2 \Rightarrow \Pr(Z \le \frac{\xi - \mu}{\sigma}) = \alpha/2,
$$

$$
\Pr(X > \zeta) = \alpha/2 \Rightarrow \Pr(Z > \frac{\zeta - \mu}{\sigma}) = \alpha/2,
$$

which implies that $-z_{\alpha/2}=\frac{\xi-\mu}{\sigma}$ $\frac{-\mu}{\sigma}$ and $z_{\alpha/2} = \frac{\zeta - \mu}{\sigma}$ $\frac{-\mu}{\sigma}$. • The $1 - \alpha$ confidence interval for X is then $[\mu - z_{\alpha/2}\sigma, \mu + z_{\alpha/2}\sigma]$.

Example 5

Suppose that $X \sim {\sf N}(\mu,\sigma^2)$ with $\mu=10$ and $\sigma=25.$ Since $z_{2.5\%}=1.96,$ the 95% confidence interval of X is:

$$
[10-1.96(25), 10+1.96(25)] = [-39, 59].
$$

The Lognormal Distribution

- If $X \sim N(\mu, \sigma^2)$ then $Y = e^X$ is said to be log-normally distributed with the same parameters.
- The distribution function of the lognormal distribution can be obtained as follows:

$$
Pr(Y \le y) = Pr(X \le ln(y))
$$

=
$$
\int_{-\infty}^{\ln(y)} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx
$$

We can then define $z = e^x \Leftrightarrow x = \ln(z) \Rightarrow dx = (1/z)dz$ and

$$
\Pr(Y \leq y) = \int_{-\infty}^{y} \frac{1}{z\sqrt{2\pi\sigma^2}} e^{-\frac{(\ln(z)-\mu)^2}{2\sigma^2}} dz
$$

Probability Density Function Comparison

- Unlike the normal density, the lognormal density function is not symmetric around its mean.
- Normally distributed variables can take values in (−∞*,* ∞), whereas lognormally distributed variables are always positive.

Computing Probabilities for Lognormal RVs

We can use the fact that the logarithm of a lognormal random variable is normally distributed to compute cumulative probabilities.

Example 6

Let $Y=e^{4+1.5Z}$ where $Z\sim {\sf N}(0,1).$ What is the probability that $Y\le 100?$

$$
Pr(Y \le 100) = Pr(e^{X} \le 100)
$$

= Pr(X \le ln(100))
= Pr(Z \le \frac{ln(100) - 4}{1.5})
= \Phi(0.4034)
= 0.6567

Confidence Interval for a Lognormal RV

Let $Y = e^{\mu + \sigma Z}$ where $Z \sim {\sf N}(0,1)$. We have that:

$$
-z_{\alpha/2} < Z \le z_{\alpha/2} \Rightarrow \mu - \sigma z_{\alpha/2} < \mu + \sigma Z \le \mu + \sigma z_{\alpha/2}
$$

$$
\Rightarrow e^{\mu - \sigma z_{\alpha/2}} < e^{\mu + \sigma Z} \le e^{\mu + \sigma z_{\alpha/2}}
$$

The $1 - \alpha$ confidence interval for Y is $[e^{\mu - \sigma z_{\alpha/2}}, e^{\mu + \sigma z_{\alpha/2}}]$.

Example 7

Let $Y = e^{4+1.5Z}$ where $Z \sim \mathsf{N}(0,1)$. The 95% confidence interval for Y is: $[e^{4-1.96(1.5)}, e^{4+1.96(1.5)}] = [2.89, 1032.71].$

Moments of a Lognormal Random Variable

Let $Y = e^{\mu + \sigma Z}$ where $Z \sim {\sf N}(0,1)$. We have that:

$$
E(Y) = e^{\mu + 0.5\sigma^2} \tag{1}
$$

$$
V(Y) = e^{2\mu + \sigma^2} (e^{\sigma^2} - 1)
$$
 (2)

$$
SD(Y) = E(Y)\sqrt{e^{\sigma^2} - 1}
$$
 (3)

Example 8

Let $Y = e^{4+1.5Z}$ where $Z \sim N(0, 1)$. The expectation and standard deviation of Y are:

$$
E(Y) = e^{4+0.5(1.5^2)} = 168.17
$$

SD(Y) = 168.17 $\sqrt{e^{1.5^2} - 1} = 489.95$

Partial Expectations

- When pricing a call option, the payoff is positive if the option is in-the-money and zero otherwise.
- We usually use an indicator function to quantify this behavior:

$$
\mathbb{1}_{\{Y>K\}} = \begin{cases} 0 & \text{if } Y \leq K \\ 1 & \text{if } Y > K \end{cases}
$$

Let $Y=e^X$ where $X\sim \mathsf{N}(\mu,\sigma^2).$ Then we have that:

$$
E(Y1_{\{Y>K\}}) = e^{\mu + \frac{1}{2}\sigma^2} \Phi\left(\frac{\mu + \sigma^2 - \ln(K)}{\sigma}\right)
$$
(4)

$$
E(K1_{\{Y>K\}}) = K \Phi\left(\frac{\mu - \ln(K)}{\sigma}\right)
$$
(5)