Statistics Preliminaries

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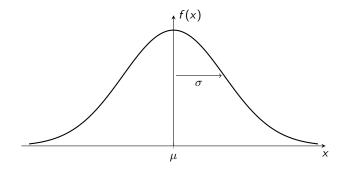
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The Normal Distribution

• We say that a random variable (RV) X is normally distributed with mean μ and standard deviation σ if its probability density function is:

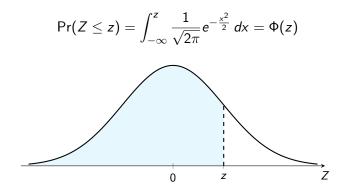
$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

and we usually write $X \sim N(\mu, \sigma^2)$.



Standard Normal Cumulative Probability

• If $Z \sim N(0, 1)$, we have that:



• Since the integral cannot be solved in closed-form, the probability must then be obtained from a table or using a computer.

Computing Normal Probabilities

 In order to compute cumulative probabilities for X ~ N(μ, σ²), we usually normalize the random variable X:

$$\Pr(X \le x) = \Pr\left(\frac{X - \mu}{\sigma} \le \frac{x - \mu}{\sigma}\right) = \Pr\left(Z \le \frac{x - \mu}{\sigma}\right) = \Phi\left(\frac{x - \mu}{\sigma}\right)$$

where $Z = \frac{X-\mu}{\sigma} \sim N(0,1)$ is called a Z-score.

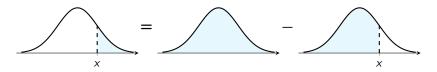
Example 1

Suppose that $X \sim N(\mu, \sigma^2)$ with $\mu = 10$ and $\sigma = 25$. What is the probability that $X \leq 0$?

$$\Pr(X \le 0) = \Pr\left(Z \le \frac{0-10}{25}\right) = \Phi(-0.40) = 0.3446.$$

Computing a Right-Tail Probability

• Right-Tail Probability: $Pr(X > x) = 1 - Pr(X \le x)$



Example 2

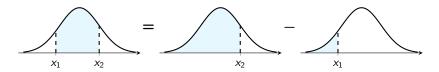
Suppose that $X \sim N(\mu, \sigma^2)$ with $\mu = 10$ and $\sigma = 25$. What is the probability that X > 12?

$$\Pr(X \le 12) = \Pr\left(Z \le \frac{12-10}{25}\right) = \Phi(0.08) = 0.5319.$$

Therefore, Pr(X > 12) = 1 - 0.5319 = 0.4681.

Computing an Interval Probability

• Interval Probability: $Pr(x_1 < X \le x_2) = Pr(X \le x_2) - Pr(X \le x_1)$



Example 3

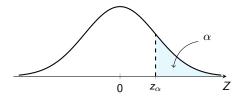
Suppose that $X \sim N(\mu, \sigma^2)$ with $\mu = 10$ and $\sigma = 25$. What is the probability that $2 < X \le 14$?

$$\begin{aligned} \Pr(X \le 14) &= \Pr\left(Z \le \frac{14-10}{25}\right) = \Phi(0.16) = 0.5636, \\ \Pr(X \le 2) &= \Pr\left(Z \le \frac{2-10}{25}\right) = \Phi(-0.32) = 0.3745. \end{aligned}$$

Therefore, $Pr(2 < X \le 14) = 0.5636 - 0.3745 = 0.1891$.

Computing a Standard Normal Right-Tail Percentile

• For a standard normal variable Z, a right-tail percentile is the value z_{α} above which we obtain a certain probability α .



• Mathematically, this means finding z_{α} such that:

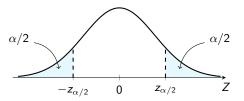
$$\Pr(Z > z_{\alpha}) = \alpha \Leftrightarrow \Pr(Z \le z_{\alpha}) = 1 - \alpha$$

• The following table shows common values for z_{α} :

α	0.05	0.025	0.01	0.005
z_{lpha}	1.64	1.96	2.33	2.58

Confidence Interval for a Standard Normal RV

• A $1 - \alpha$ confidence interval (CI) defines left and right percentiles such that the probability on each side is $\alpha/2$.



• For a standard normal variable Z, the symmetry of its PDF implies:

$$\mathsf{Pr}(Z \leq -z_{lpha/2}) = \mathsf{Pr}(Z > z_{lpha/2}) = lpha/2$$

Example 4

Since $z_{2.5\%} = 1.96$, the 95% confidence interval of Z is [-1.96, 1.96]. This means that if we randomly sample this variable 100,000 times, approximately 95,000 observations will fall inside this interval.

Confidence Interval for a Normal RV

If X ~ N(μ, σ²), its confidence interval is determined by ξ and ζ such that:

$$\Pr(X \le \xi) = \alpha/2 \Rightarrow \Pr(Z \le \frac{\xi - \mu}{\sigma}) = \alpha/2,$$

$$\Pr(X > \zeta) = \alpha/2 \Rightarrow \Pr(Z > \frac{\zeta - \mu}{\sigma}) = \alpha/2,$$

which implies that $-z_{\alpha/2} = \frac{\xi - \mu}{\sigma}$ and $z_{\alpha/2} = \frac{\zeta - \mu}{\sigma}$.

• The $1 - \alpha$ confidence interval for X is then $[\mu - z_{\alpha/2}\sigma, \mu + z_{\alpha/2}\sigma]$.

Example 5

Suppose that $X \sim N(\mu, \sigma^2)$ with $\mu = 10$ and $\sigma = 25$. Since $z_{2.5\%} = 1.96$, the 95% confidence interval of X is:

$$[10 - 1.96(25), 10 + 1.96(25)] = [-39, 59].$$

The Lognormal Distribution

- If $X \sim N(\mu, \sigma^2)$ then $Y = e^X$ is said to be log-normally distributed with the same parameters.
- The distribution function of the lognormal distribution can be obtained as follows:

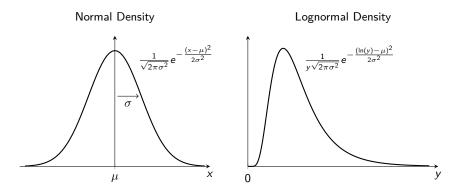
$$\Pr(Y \le y) = \Pr(X \le \ln(y))$$
$$= \int_{-\infty}^{\ln(y)} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

• We can then define $z = e^x \Leftrightarrow x = \ln(z) \Rightarrow dx = (1/z)dz$ and

$$\Pr(Y \le y) = \int_{-\infty}^{y} \frac{1}{z\sqrt{2\pi\sigma^2}} e^{-\frac{(\ln(z)-\mu)^2}{2\sigma^2}} dz$$

Probability Density Function Comparison

- Unlike the normal density, the lognormal density function is not symmetric around its mean.
- Normally distributed variables can take values in $(-\infty, \infty)$, whereas lognormally distributed variables are always positive.



Computing Probabilities for Lognormal RVs

 We can use the fact that the logarithm of a lognormal random variable is normally distributed to compute cumulative probabilities.

Example 6

Let $Y = e^{4+1.5Z}$ where $Z \sim N(0, 1)$. What is the probability that $Y \le 100$?

$$Pr(Y \le 100) = Pr(e^X \le 100)$$

= Pr(X \le ln(100))
= Pr(Z \le \frac{ln(100)-4}{1.5})
= \Phi(0.4034)
= 0.6567

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Confidence Interval for a Lognormal RV

• Let $Y = e^{\mu + \sigma Z}$ where $Z \sim N(0, 1)$. We have that:

$$\begin{aligned} -z_{\alpha/2} < Z \le z_{\alpha/2} \Rightarrow \mu - \sigma z_{\alpha/2} < \mu + \sigma Z \le \mu + \sigma z_{\alpha/2} \\ \Rightarrow e^{\mu - \sigma z_{\alpha/2}} < e^{\mu + \sigma Z} \le e^{\mu + \sigma z_{\alpha/2}} \end{aligned}$$

• The $1 - \alpha$ confidence interval for Y is $[e^{\mu - \sigma z_{\alpha/2}}, e^{\mu + \sigma z_{\alpha/2}}]$.

Example 7

Let $Y = e^{4+1.5Z}$ where $Z \sim N(0, 1)$. The 95% confidence interval for Y is:

$$[e^{4-1.96(1.5)}, e^{4+1.96(1.5)}] = [2.89, 1032.71].$$

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Moments of a Lognormal Random Variable

• Let $Y = e^{\mu + \sigma Z}$ where $Z \sim N(0, 1)$. We have that:

$$\mathsf{E}(Y) = e^{\mu + 0.5\sigma^2} \tag{1}$$

$$V(Y) = e^{2\mu + \sigma^2} (e^{\sigma^2} - 1)$$
 (2)

$$SD(Y) = E(Y)\sqrt{e^{\sigma^2} - 1}$$
(3)

Example 8

Let $Y = e^{4+1.5Z}$ where $Z \sim N(0, 1)$. The expectation and standard deviation of Y are:

$$E(Y) = e^{4+0.5(1.5^2)} = 168.17$$
$$SD(Y) = 168.17\sqrt{e^{1.5^2} - 1} = 489.95$$

Partial Expectations

- When pricing a call option, the payoff is positive if the option is in-the-money and zero otherwise.
- We usually use an indicator function to quantify this behavior:

$$\mathbb{1}_{\{Y>K\}} = \begin{cases} 0 & \text{if } Y \leq K \\ 1 & \text{if } Y > K \end{cases}$$

• Let $Y = e^X$ where $X \sim N(\mu, \sigma^2)$. Then we have that:

$$E\left(Y\mathbb{1}_{\{Y>K\}}\right) = e^{\mu + \frac{1}{2}\sigma^2} \Phi\left(\frac{\mu + \sigma^2 - \ln(K)}{\sigma}\right)$$
(4)
$$E\left(K\mathbb{1}_{\{Y>K\}}\right) = K \Phi\left(\frac{\mu - \ln(K)}{\sigma}\right)$$
(5)